



On-line evaluation of mono-exponential decays

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This Note describes the algorithm used within AcqNmr for three-parameter mono-exponential fits noisy decay functions sampled at N values of the timing parameter τ .

Let $[m_k]$ and $[\tau_k]$ be, respectively, the arrays of average data-window magnitudes and of the arrayed parameter values (the index k ranges over all considered blocks, assumed to be n in number, $n \geq 3$). These data are to be fitted by the theoretical formula (the hypothesis)

$$(1) \quad m_k(\tau_k) = a + b e^{-r\tau_k}$$

where a, b , and r are some as yet unknown parameters. This requires a non-linear, three-parameter, least-squares fit in which one minimizes the quantity

$$(2) \quad Q(a, b, r) = \sum_k [m_k(\tau_k) - (a + b e^{-r\tau_k})]^2$$

with respect to a , b and r .

It is convenient to split the task into two distinct parts.

1. Assuming the value of r to be fixed, the formula is linear with respect to a and b . The optimal values of these two parameters are therefore easily determined using standard linear-correlation formulae [1,2,3]. The resulting 'optimal' values of a and b , denoted as a_1 , b_1 , and the corresponding value of Q and Q_1 thus become non-linear functions of r , i.e., $a_1 = a_1(r)$, $b_1 = b_1(r)$ and $Q_1 = Q_1(r)$, with $Q_1(r)$ being approximately quadratic around its absolute minimum.
2. Using the function which calculates $Q_1(r)$ for any value of r , its minimum value is then determined numerically using the standard Brent's algorithm [4]. Clearly, if the minimum of $Q_1(r)$ occurs at $r = r_2$ then $Q_2 = Q_1(r_2) = Q(a_1(r_2), b_1(r_2), r_2)$ coincides with the absolute minimum of $Q(a, b, r)$.

What we have gained in addition to the optimal fit is the possibility to evaluate the function $Q_1(r)$ for any r in the vicinity of the optimum at r_2 , where we expect it to be approximately quadratic with the quadratic coefficient related to the *confidence interval* of r . Notice that, along the curve $Q_1(r) = Q(a_1(r), b_1(r), r)$, the parameters a, b are dynamically varied, keeping them optimal for every single value of r . This is essential since otherwise the error estimates for r would be grossly over-optimistic.

Numeric values of the *confidence interval* are based on the least significant increment of Q . Assuming that the optimum value Q_2 of Q is due entirely to random experimental errors (this, of course, is false for non-exponential decays), the least significant increment $\Delta_\alpha Q$ can be determined for any given significance level α by means of the *Fisher statistics* [2,3] with both degrees of freedom set to $n-1$. The confidence interval $\Delta_\alpha r$ for r then comprises the r values for which $Q_1(r) - Q_2 \leq \Delta_\alpha Q$ and the *probable error* $e = \Delta_\alpha r / 2$ is obtained, as usual, by setting $\alpha = 0.69\dots$. This may sound complicated but it actually turns out that, in the quadratic case, the result is excellently approximated by the simple formula:

$$(3) \quad e = \sqrt{\frac{1}{(n-1)} \frac{Q_2}{Q_1''(r_2)}}$$

where $Q_1''(r)$ is the second derivative of $Q_1(r)$ which is easily estimated numerically by standard methods.

References:

1. Radhakrishna R., "Linear Statistical Inference and Its Applications", John Wiley, New York, 1973.
2. Feller W., "An Introduction to Probability Theory and Its Applications", John Wiley, New York, 1966.
3. Cramer H., "Mathematical Methods of Statistics", Princeton University Press, Princeton, NJ, 1946.
4. Press W.H., Teukolsky S.A., Vetterling W.T., Flannery B.P., Numerical Recipes in C, "The Art of Scientific Computing", Cambridge University Press, 1992.