



On-line evaluation of mono-exponential decays

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This Note describes the algorithm used within AcqNmr for three-parameter mono-exponential fits noisy decay functions sampled at N values of the timing parameter τ .

Let $[m_k]$ and $[\tau_k]$ be, respectively, the arrays of average data-window magnitudes and of the arrayed parameter values (the index k ranges over all considered blocks, assumed to be n in number, $n \geq 3$). These data are to be fitted by the theoretical formula (the hypothesis)

$$(1) \quad m_k(\tau_k) = a + b e^{-r\tau_k}$$

where a, b , and r are some as yet unknown parameters. This requires a non-linear, three-parameter, least-squares fit in which one minimizes the quantity

$$(2) \quad Q(a, b, r) = \sum_k [m_k(\tau_k) - (a + b e^{-r\tau_k})]^2$$

with respect to a , b and r .

It is convenient to split the task into two distinct parts.

1. Assuming the value of r to be fixed, the formula is linear with respect to a and b . The optimal values of these two parameters are therefore easily determined using standard linear-correlation formulae [1,2,3]. The resulting 'optimal' values of a and b , denoted as a_1 , b_1 , and the corresponding value of Q and Q_1 thus become non-linear functions of r , i.e., $a_1 = a_1(r)$, $b_1 = b_1(r)$ and $Q_1 = Q_1(r)$, with $Q_1(r)$ being approximately quadratic around its absolute minimum.
2. Using the function which calculates $Q_1(r)$ for any value of r , its minimum value is then determined numerically using the standard Brent's algorithm [4]. Clearly, if the minimum of $Q_1(r)$ occurs at $r = r_2$ then $Q_2 = Q_1(r_2) = Q(a_1(r_2), b_1(r_2), r_2)$ coincides with the absolute minimum of $Q(a, b, r)$.

What we have gained in addition to the optimal fit is the possibility to evaluate the function $Q_1(r)$ for any r in the vicinity of the optimum at r_2 , where we expect it to be approximately quadratic with the quadratic coefficient related to the *confidence interval* of r . Notice that, along the curve $Q_1(r) = Q(a_1(r), b_1(r), r)$, the parameters a, b are dynamically varied, keeping them optimal for every single value of r . This is essential since otherwise the error estimates for r would be grossly over-optimistic.

Numeric values of the *confidence interval* are based on the least significant increment of Q . Assuming that the optimum value Q_2 of Q is due entirely to random experimental errors (this, of course, is false for non-exponential decays), the least significant increment $\Delta_\alpha Q$ can be determined for any given significance level α by means of the *Fisher statistics* [2,3] with both degrees of freedom set to $n-1$. The confidence interval $\Delta_\alpha r$ for r then comprises the r values for which $Q_1(r) - Q_2 \leq \Delta_\alpha Q$ and the *probable error* $e = \Delta_\alpha r / 2$ is obtained, as usual, by setting $\alpha = 0.69\dots$. This may sound complicated but it actually turns out that, in the quadratic case, the result is excellently approximated by the simple formula:

$$(3) \quad e = \sqrt{\frac{1}{(n-1)} \frac{Q_2}{Q_1''(r_2)}}$$

where $Q_1''(r)$ is the second derivative of $Q_1(r)$ which is easily estimated numerically by standard methods.

References:

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